

On sampling determinantal point processes

Guillaume Gautier

Ph.D. defense

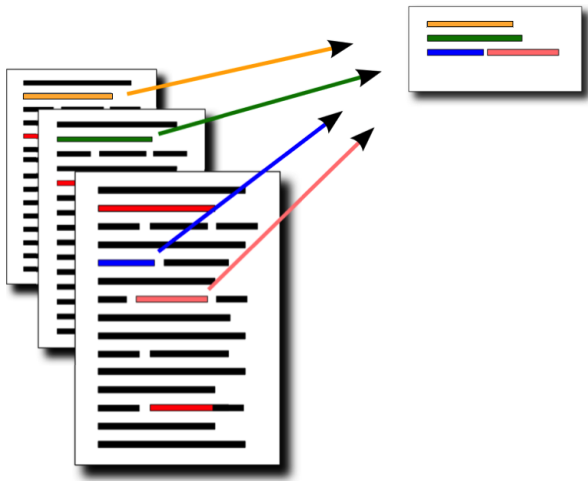


Advisors: Rémi Bardenet, Michal Valko

May 19, 2020

Text summarization

Extract **diverse** sentences of a large corpus to build a representative summary.



Recommendation systems

Two possible sets of answers of an image search engine to the query “bolt”.

relevance



relevance
+
diversity

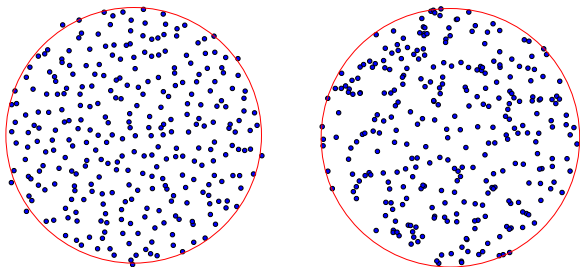


- ▶ Use DPPs to enforce diversity among the recommended items.

Numerical integration

Use random **repulsive** points as quadrature nodes to estimate an integral

$$\int f(x)\mu(dx) \approx \sum_{n=1}^N \omega_n f(x_n).$$



Bardenet and Hardy (2016, 2020)

- ▶ Prove faster rate of convergence with DPP points than i.i.d. points.
- ▶ Efficient sampler to put theory into practice?

My Ph.D. in a nutshell

Finite projection DPP

Approximate sampling
Linear programming



ICML, 2017

β -ensembles

Gibbs sampling
Random matrices

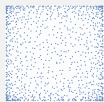
eigvals $\begin{bmatrix} \text{red diagonal} & 0 \\ \text{blue diagonal} & \text{red diagonal} \\ 0 & \text{blue diagonal} \end{bmatrix}$

Submitted, 2020

DPP sampling

Monte Carlo integration

Exact sampling
Random linear system



NeurIPS, 2019

Python toolbox

DPPy  
Reproducible research

```
from dppy import *  
# [...]  
dpp.sample()
```

JMLR-MLOSS, 2019

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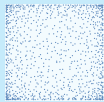
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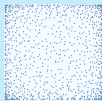
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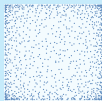
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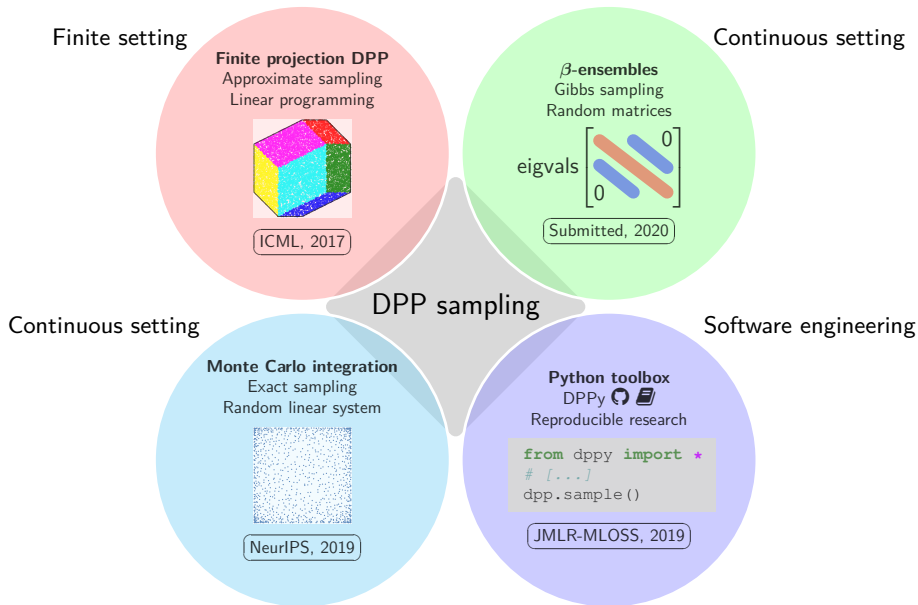
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Focus of the presentation

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ICML, 2017



Continuous setting

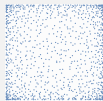
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Overview

Introduction

DPP basics

- Some insights on finite DPPs

- Finite projection DPPs

- Continuous projection DPPs

- Exact sampling from finite projection DPPs

- Approximate sampling from finite projection DPPs

Contributions

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Conclusion

- Summary of contributions

- Open questions and perspectives

Some insights on finite DPPs

► Point process



Some insights on finite DPPs

- ▶ Point process

$$\mathcal{X} \subset \left\{ \begin{array}{c} \text{[Jamaican Bobsled Team runner]} \\ \text{[Jamaican Bobsled Team runner]} \\ \text{[Chihuahua dog]} \end{array} , \begin{array}{c} \text{[Chevrolet Spark car]} \\ \text{[Yellow bolt]} \end{array} , \dots , \begin{array}{c} \text{[Chihuahua dog]} \end{array} \right\}$$

- ▶ Diversity

$$\mathbb{P}[\{\text{[Chevrolet Spark car]}, \text{[Jamaican Bobsled Team runner]}\} \subset \mathcal{X}] \geq \mathbb{P}[\{\text{[Chevrolet Spark car]}, \text{[Chihuahua dog]}\} \subset \mathcal{X}]$$

Some insights on finite DPPs

- ▶ Point process

$$\mathcal{X} \subset \left\{ \begin{array}{c} \text{[Jamaican athlete]} \\ \text{[Jamaican athlete]} \\ \text{[Duke the dog]} \end{array} , \begin{array}{c} \text{[Chevrolet Volt]} \\ \text{[Bolt bolt]} \end{array} , \dots , \begin{array}{c} \text{[Chihuahua]} \end{array} \right\}$$

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- ▶ Similarity matrix

$$\mathbf{K} \begin{array}{c} \text{[Chevrolet Volt]} \\ \text{[Jamaican athlete]} \end{array}$$

Some insights on finite DPPs

- ▶ Point process

$$\mathcal{X} \subset \left\{ \begin{array}{c} \text{[Jamaican Bobsled Team runner]} \\ \text{[Jamaican Bobsled Team runner]} \\ \text{[Dog head]} \\ \text{[Chevrolet Bolt EV car]} \\ \text{[Bolt nut]} \\ \dots \\ \text{[Chihuahua dog]} \end{array} \right\}$$

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- ▶ Inclusion probabilities

$$\mathbb{P}[\{\text{[Chevrolet Bolt EV car]}, \text{[Jamaican Bobsled Team runner]}\} \subset \mathcal{X}] = \det \begin{bmatrix} \mathbf{K} \begin{array}{c} \text{[Chevrolet Bolt EV car]} \\ \text{[Chevrolet Bolt EV car]} \end{array} & \mathbf{K} \begin{array}{c} \text{[Chevrolet Bolt EV car]} \\ \text{[Jamaican Bobsled Team runner]} \end{array} \\ \mathbf{K} \begin{array}{c} \text{[Jamaican Bobsled Team runner]} \\ \text{[Chevrolet Bolt EV car]} \end{array} & \mathbf{K} \begin{array}{c} \text{[Jamaican Bobsled Team runner]} \\ \text{[Jamaican Bobsled Team runner]} \end{array} \end{bmatrix}$$

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- ▶ Sufficient conditions for existence

$$\mathbf{K}^T = \mathbf{K} \quad \text{and} \quad 0 \preceq \mathbf{K} \preceq I.$$

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Finite projection DPP

Definition

Consider $\mathbf{K} \in \mathbb{R}^{M \times M}$ such that $\mathbf{K}^\top = \mathbf{K}$ and $\mathbf{K}^2 = \mathbf{K}$.

The point process \mathcal{X} defined by

$$\mathbb{P}[S \subset \mathcal{X}] = \det \mathbf{K}_S, \quad \forall S \subset \{1, \dots, M\},$$

is called a projection DPP with kernel \mathbf{K} .

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$$\mathbb{P}[\mathcal{X} = B] = \det \mathbf{K}_B \mathbf{1}_{|B|=N}.$$

Finite projection DPP

Example

Consider the $N \times M$ feature matrix $\Phi = [\phi_1, \dots, \phi_M]$, such that rank $\Phi = N$, and build the kernel

$$\mathbf{K} = \Phi^\top [\Phi \Phi^\top]^{-1} \Phi.$$

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Example

Consider the $N \times M$ feature matrix $\Phi = [\phi_1, \dots, \phi_M]$, such that $\text{rank } \Phi = N$, and build the kernel

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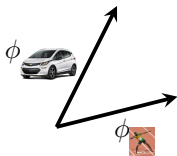
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$$\mathbb{P}[\mathcal{X} = \{\text{car}, \text{tree}\}] \propto \text{volume}^2$$



Finite projection DPP

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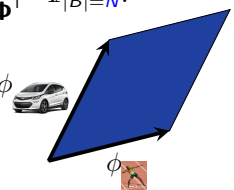
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$$\mathbb{P}[\mathcal{X} = \{\text{car}, \text{tree}\}] \propto \text{volume}^2 \phi_{\text{car}} \phi_{\text{tree}}$$


- ▶ The support is formed by collections of columns of Φ forming a basis of \mathbb{R}^N ,
 $\mathcal{B} \triangleq \{B; |B| = N, \text{ and } \det \Phi_{:B} \neq 0\}.$

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Continuous projection DPP

Definition

Let $\phi_0, \dots, \phi_{N-1}$ be orthonormal functions in $L^2(\mathbb{X}, \mu)$ and

$$K(x, y) = \sum_{k=0}^{N-1} \phi_k(x) \phi_k(y).$$

Take (x_1, \dots, x_N) with joint probability distribution

$$\frac{1}{N!} \det[K(x_i, x_j)]_{i,j=1}^N \prod_{n=1}^N \mu(dx_n).$$

Then $\mathcal{X} \triangleq \{x_1, \dots, x_N\} \subset \mathbb{X}$ defines a projection DPP with kernel K .

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Considering

$$\blacktriangleright \mathbb{X} = \{1, \dots, M\},$$

$$\blacktriangleright \mu = \sum_{m=1}^M \delta_m,$$

one recovers the finite case with $\mathbf{K} = \Phi^T \Phi$ and $\Phi \Phi^T = I_N$.

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The goal is to generate a random **subset** \mathcal{X} , such that

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Consider the eigendecomposition $\mathbf{K} = \mathbf{\Phi}^\top \mathbf{\Phi}$, where $\mathbf{\Phi} \mathbf{\Phi}^\top = I_N$.

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$$\mathbb{P}[(x_1, \dots, x_n)] = \frac{1}{N!} \text{volume}^2\{\phi_{x_1}, \dots, \phi_{x_N}\} = \frac{1}{N!} \det \mathbf{K}_{\{x_1, \dots, x_N\}}.$$

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- ▶ The procedure is akin to Gram-Schmidt orthogonalization $\mathcal{O}(MN^2)$.

Illustration of the chain rule ($M = 24, N = 2$)

Text to summarize using $N = 2$ sentences.

But a dream within a dream?
 Is all that we see or seem
 One from the pitiless wave?
 O God! can I not save
 Them with a tighter clasp?
 O God! can I not grasp
 While I weep--while I weep!
 Through my fingers to the deep,
 How few! yet how they creep
 Grains of the golden sand--
 And I hold within my hand
 Of a surf-tormented shore,
 I stand amid the roar
 Is but a dream within a dream.
 All that we see or seem
 Is it therefore the less gone?
 In a vision, or in none,
 In a night, or in a day,
 Yet if hope has flown away
 That my days have been a dream;
 You are not wrong, who deem
 Thus much let me avow--
 And, in parting from you now,
 Take this kiss upon the brow!

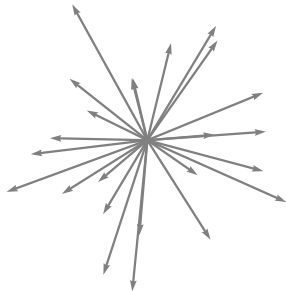
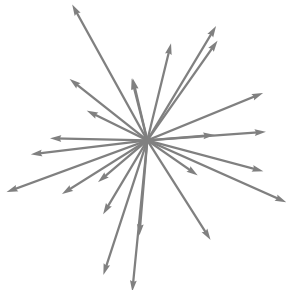


Illustration of the chain rule ($M = 24, N = 2$)

Select the first sentence,



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I stand amid the roar
Is but a dream within a dream.
All that we see or seem
Is it therefore the less gone?
In a vision, or in none,
In a night, or in a day,
Yet if hope has flown away
That my days have been a dream;
You are not wrong, who deem
Thus much let me avow--
And, in parting from you now,
Take this kiss upon the brow!



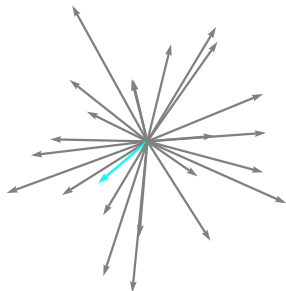
$$\mathbb{P}[x_1 = x] = \frac{1}{2} \|\phi_x\|^2.$$

Illustration of the chain rule ($M = 24, N = 2$)

Select the first sentence,



But a dream within a dream?
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O God! can I not save
Them with a tighter clasp?
O God! can I not grasp
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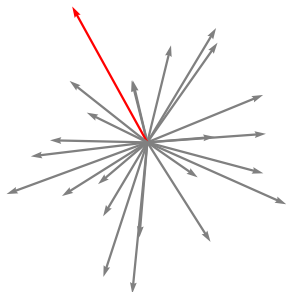


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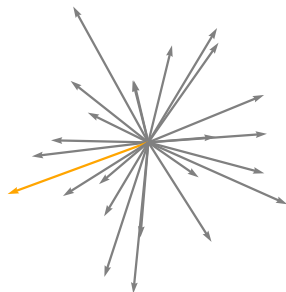


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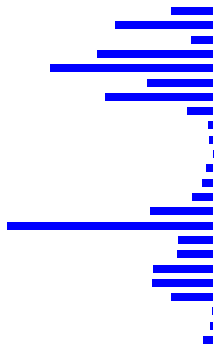
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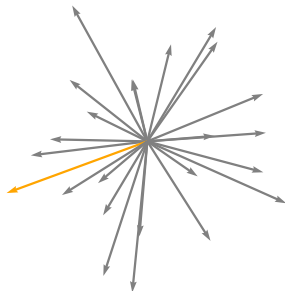
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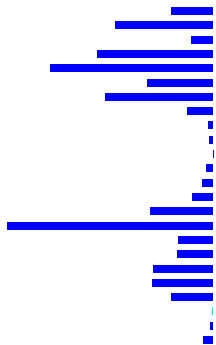
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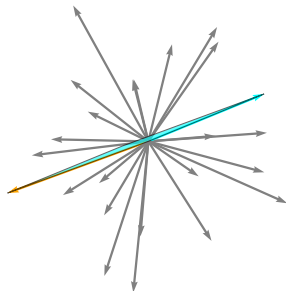
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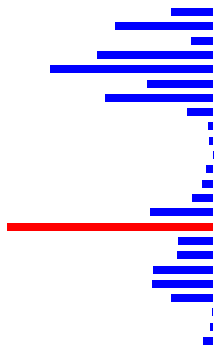
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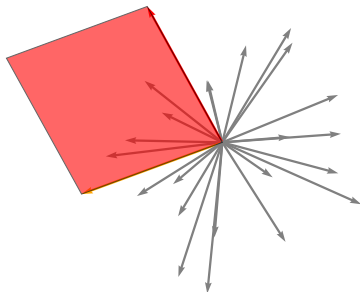
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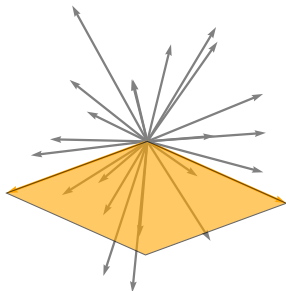
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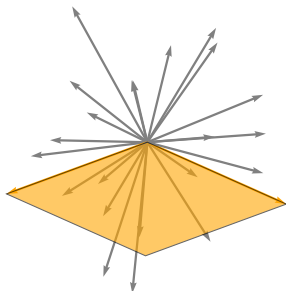


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The basis-exchange walk

Metropolis Hastings kernel

(Feder and Mihail, 1992; Anari, Gharan, and Rezaei, 2016; Li, Jegelka, and Sra, 2016; Hermon and Salez, 2019)

- ▶ Starting from $B_0 \in \mathcal{B}$.

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Illustration of the basis-exchange walk ($M = 24, N = 7$)

$$B_0 = \{1, 3, 9, 12, 13, 18, 24\}$$

But a dream within a dream?

Is all that we see or seem

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O God! can I not save

Them with a tighter clasp?

O God! can I not grasp

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Conceptual shift: sampling by solving randomized linear programs

Finite projection DPP

Approximate sampling
Linear programming



ICML, 2017

β -ensembles

Gibbs sampling
Random matrices

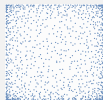
eigvals $\begin{bmatrix} \text{red diagonal} & 0 \\ \text{blue diagonal} & \text{red diagonal} \\ 0 & \text{blue diagonal} \end{bmatrix}$

Submitted, 2020

DPP sampling

Monte Carlo integration

Exact sampling
Random linear system



NeurIPS, 2019

Python toolbox

DPPy  
Reproducible research

```
from dppy import *  
# [...]  
dpp.sample()
```

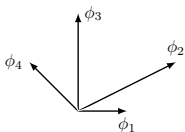
JMLR-MLOSS, 2019

Continuous embedding of the support

The support of finite projection DPPs, characterized by

$$\mathcal{B} \triangleq \{B; |B| = N, \text{ and } \det \Phi_{:B} \neq 0\},$$

has the following geometrical representation.

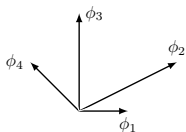
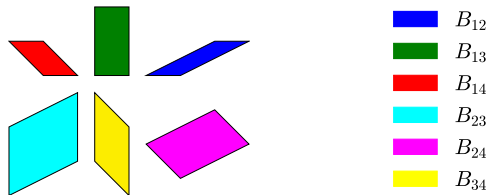


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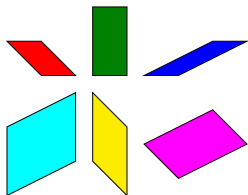
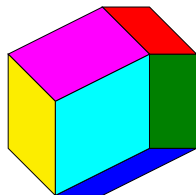


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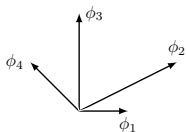
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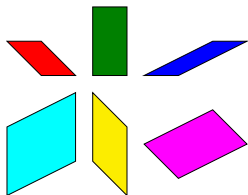
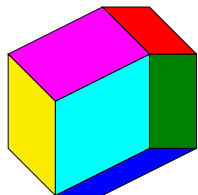


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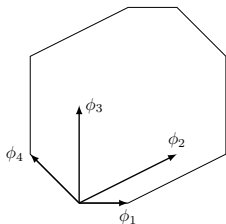
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How to identify the tiles?

Tiling of a zonotope (Dyer and Frieze, 1994)

Definition (Zonotope)

Let $\Phi \in \mathbb{R}^{N \times M}$ such that rank $\Phi = N$

$$\mathcal{Z}(\Phi) \triangleq \Phi[0, 1]^M.$$

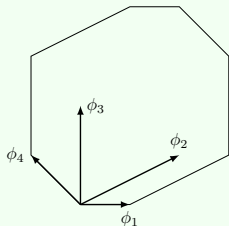
Tiling of a zonotope (Dyer and Frieze, 1994)

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Let $\Phi \in \mathbb{R}^{N \times M}$ such that $\text{rank } \Phi = N$

$$\mathcal{Z}(\Phi) \triangleq \Phi[0, 1]^M.$$

Example ($M = 4, N = 2$)



Tiling of a zonotope (Dyer and Frieze, 1994)

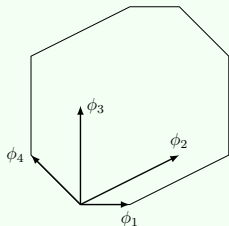
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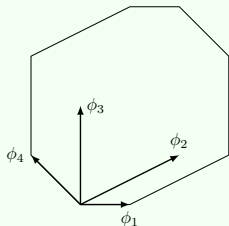
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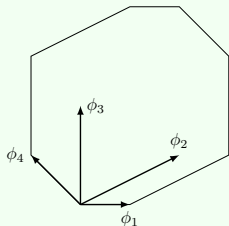
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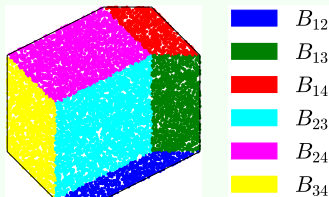
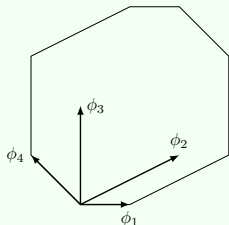
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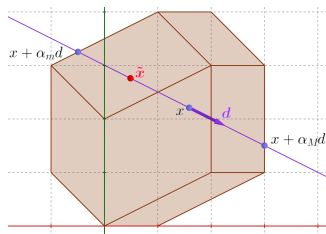
Random walk on zonotope $\xrightarrow{(LP)}$ random walk on tiles

Gautier, Bardenet, and Valko (2017)

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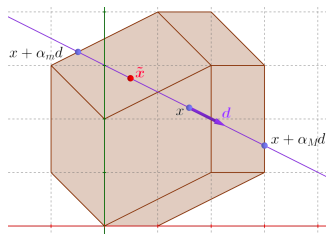


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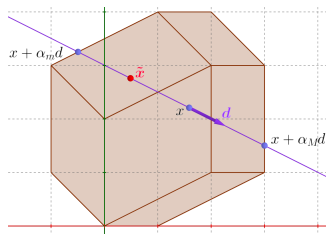
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► Target density on $\mathcal{Z}(\Phi)$

$$\pi(x) = \sum_{B \in \mathcal{B}} C_B \times \mathbb{1}_{\mathcal{Z}(\Phi; B)}(x).$$

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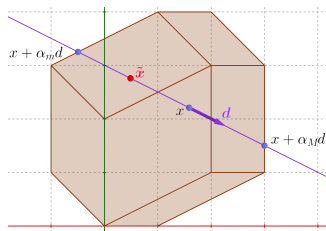
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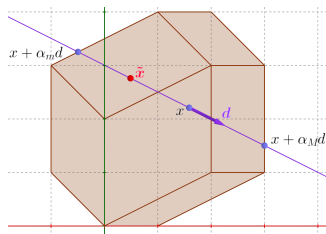
► Limiting distribution on \mathcal{B}

$$\mathbb{P}[B_x = B] = C_B \times |\det \Phi_{:B}|.$$

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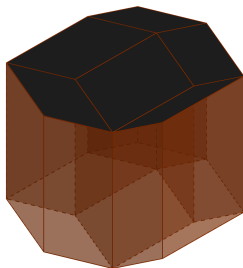
$$\mathbb{P}[B_x = B] = C_B \times |\det \Phi_{:B}|.$$

How to make $\mathbb{P}[B_x = B] \propto (\det \Phi_{:B})^2$?

Hit-and-run with acceptance ratio = 1

The target density on $\mathcal{Z}(\Phi)$ is uniform,

$$\pi(x) \propto \sum_{B \in \mathcal{B}} 1 \times \mathbb{1}_{\mathcal{Z}(\Phi:B)}(x).$$



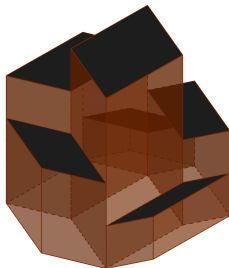
The limiting distribution on \mathcal{B} takes the form

$$\mathbb{P}[B_x = B] \propto 1 \times |\det \Phi_{:B}| = |\det \Phi_{:B}|^1.$$

Hit-and-run with acceptance ratio = $\left| \frac{\det \Phi_{:\tilde{B}}}{\det \Phi_{:B}} \right|$

The target density on $\mathcal{Z}(\Phi)$ is given by

$$\pi(x) \propto \sum_{B \in \mathcal{B}} |\det \Phi_{:B}| \times \mathbb{1}_{\mathcal{Z}(\Phi_{:B})}(x).$$



The limiting distribution on \mathcal{B} takes the form

$$\mathbb{P}[B_x = B] \propto |\det \Phi_{:B}| \times |\det \Phi_{:B}| = (\det \Phi_{:B})^2.$$

Illustration of the zonotope walk ($M = 24, N = 7$)

$$B_0 = \{1, 3, 9, 12, 13, 18, 24\}$$

But a dream within a dream?

Is all that we see or seem

One from the pitiless wave?

O God! can I not save

Them with a tighter clasp?

O God! can I not grasp

While I weep--while I weep!

Through my fingers to the deep,

How few! yet how they creep

Grains of the golden sand--

And I hold within my hand

Of a surf-tormented shore,

I stand amid the roar

Is but a dream within a dream.

All that we see or seem

Is it therefore the less gone?

In a vision, or in none,

In a night, or in a day,

Yet if hope has flown away

That my days have been a dream;

You are not wrong, who deem

Thus much let me avow--

And, in parting from you now,

Take this kiss upon the brow!

Illustration of the zonotope walk ($M = 24, N = 7$)

$$B_0 = \{1, 3, 9, 12, 13, 18, 24\} \quad B_1 = (B_0 \setminus \{1, 9, 13, 18\}) \\ \cup \{6, 8, 14, 17\}$$

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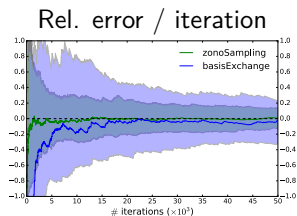
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Comparison of the zonotope and basis-exchange walks

Relative error of the estimation of $\mathbb{P}[\{x_1, x_2, x_3\} \subset \mathcal{X}] = \det \mathbf{K}_{\{x_1, x_2, x_3\}}$.

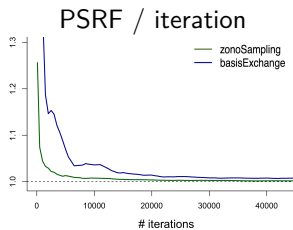
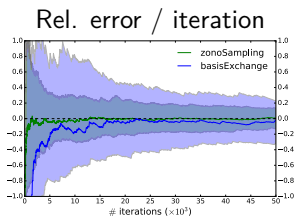
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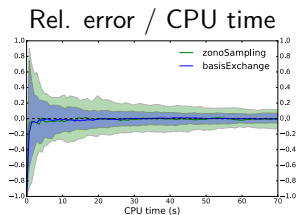
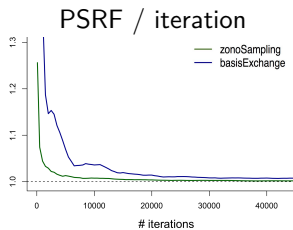
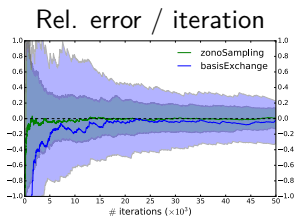
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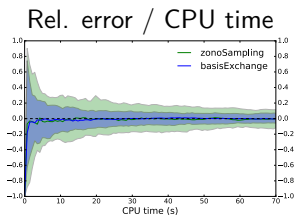
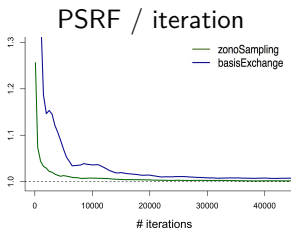
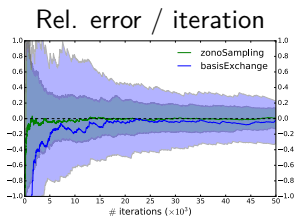
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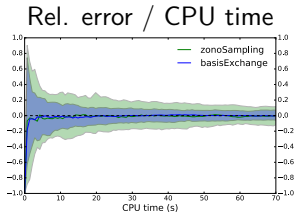
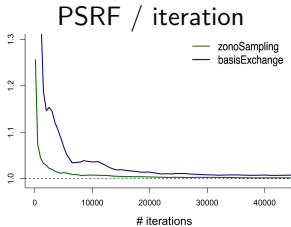
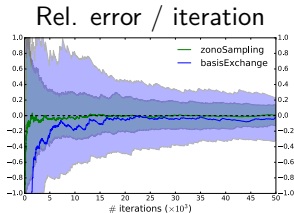
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	basis-exchange		zonotope	
Exploration of the support	✓	(lazy)	✓✓✓	
Empirical mixing	✓✓		✓✓✓	
Cost per iteration	✓✓✓	det	✓	det + 3 LPs
Theoretical guarantees	✓✓	$\text{poly}(M, N)$?	$\text{poly}(M, N)$?

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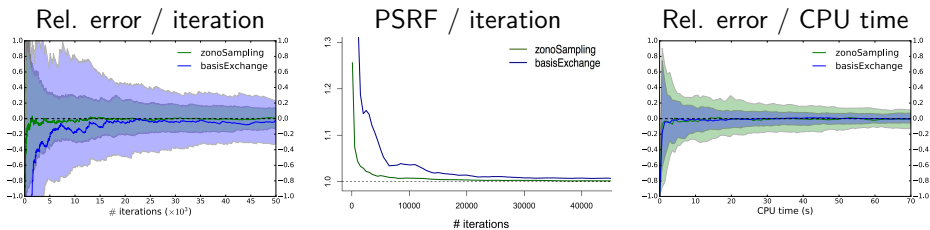


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The zonotope walk is sample efficient.

Comparison of the zonotope and basis-exchange walks

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The zonotope walk is sample efficient.

Can we generalize the idea to the continuous setting?

Overview

Introduction

DPP basics

Some insights on finite DPPs

Finite projection DPPs

Continuous projection DPPs

Exact sampling from finite projection DPPs

Approximate sampling from finite projection DPPs

Contributions

Zonotope sampling for finite projection DPPs

Transition

Fast sampling from β -ensembles

Conclusion

Summary of contributions

Open questions and perspectives

Sampling by solving randomized linear programs?

Finite case

Linear Programming (LP)

$$\begin{aligned}
 \min_y \quad & c^\top y \\
 \text{s.t.} \quad & \varphi_1^\top y = x_1 \\
 & \vdots \\
 & \varphi_N^\top y = x_N \\
 & 0 \leq y \leq 1
 \end{aligned}$$

- ▶ Unique solution y^* 😊
- ▶ Efficient solvers 😊
- ▶ “Support” of the solution
 - ▶ $|i ; 0 < y_i^* < 1| = N$ 😊

Sampling by solving randomized linear programs?

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Continuous case (dimension d)

Linear Semi Infinite Programming (LSIP)

$$\begin{aligned} \min_\nu \quad & \int c(x) \nu(dx) \\ \text{s.t.} \quad & \int \varphi_1(x) \nu(dx) = m_1 \\ & \vdots \\ & \int \varphi_N(x) \nu(dx) = m_N \\ & \text{“} 0 \leq \mu \leq 1 \text{”} \end{aligned}$$

- ▶ No unique solution 😞
- ▶ No efficient solvers 😞
- ▶ Structure of the support of solutions
 - ▶ $\exists \nu^*$ s.t. $|\text{supp } \nu^*| \leq N$.

(Goberna and López, 2014)

Sampling by solving randomized linear programs?

Dimension $d > 1$

For polynomials functions c and φ_n

$$\begin{aligned} \min_{\nu} \quad & \int c(x)\nu(dx) \\ \text{s.t.} \quad & \int \varphi_1(x)\nu(dx) = m_1 \\ & \quad \quad \quad \vdots \\ & \int \varphi_N(x)\nu(dx) = m_N \end{aligned}$$

- ▶ No unique solution ☹️
- ▶ Efficient solvers? (Lasserre, 2010)
 - ▶ hierarchy of SDP relaxations 😊
 - ▶ works for small d and N ☹️
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 - ▶ $\exists \nu^*$ s.t. $|\text{supp } \nu^*| \leq N$
 - ▶ unstable support extraction ☹️

Sampling by solving randomized linear programs?

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$$\min_{\nu} \int c(x)\nu(dx)$$

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Dimension $d = 1$

Truncated moment problem

$$\min_{\nu} \mathbb{E}_{\nu}[X^{2N}]$$

$$\text{s.t.} \quad \mathbb{E}_{\nu}[X] = m_1$$

$$\vdots$$

$$\mathbb{E}_{\nu}[X^{2N-1}] = m_{2N-1}$$

- ▶ Unique solution $\nu^* = \sum_{n=1}^N \omega_n \delta_{x_n}$ 😊
- ▶ Unstable support extraction ☹️

Sampling by solving randomized linear programs?

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Linear Programming (LP)

$$\begin{aligned} \min_y \quad & c^\top y \\ \text{s.t.} \quad & \varphi_1^\top y = x_1 \\ & \vdots \\ & \varphi_N^\top y = x_N \\ & 0 \leq y \leq 1 \end{aligned}$$

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Truncated moment problem

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- ▶ Unique solution $\nu^* = \sum_{n=1}^N \omega_n \delta_{x_n}$ 😊
- ▶ Unstable support extraction ☹️

How to randomize the moment constraints s.t. $\{x_1, \dots, x_N\} \sim$ target DPP?

Sampling by solving randomized linear programs?

Finite case

Linear Programming (LP)

$$\begin{aligned} \min_y \quad & c^\top y \\ \text{s.t.} \quad & \varphi_1^\top y = x_1 \\ & \vdots \\ & \varphi_N^\top y = x_N \\ & 0 \leq y \leq 1 \end{aligned}$$

- ▶ Unique solution y^* 😊
- ▶ Efficient solvers 😊
- ▶ “Support” of the solution
 - ▶ $|i; 0 < y_i^* < 1| = N$ 😊

Dimension $d = 1$

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$(\omega_n), (x_n)$ define a quadrature rule
(**RyBo15**; Dette and Studden, 1997)

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Sampling by computing the eigenvalues of random tridiagonal matrices

Dimension $d = 1$

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Reparametrize ν^* via the 3-terms recurrence relation \perp polynomials encoded by

$$J_{\mathbf{a}, \mathbf{b}} \triangleq \begin{bmatrix} a_1 & \sqrt{b_1} & & & (0) \\ \sqrt{b_1} & a_2 & & & \\ & & \ddots & & \\ & & & \ddots & \\ (0) & & & \sqrt{b_{N-1}} & a_N \end{bmatrix}$$

- ▶ $\{x_1, \dots, x_N\} = \text{eigvals } J_{\mathbf{a}, \mathbf{b}}$
- ▶ Computational cost $\mathcal{O}(N^2)$

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Sampling 1D continuous projection DPPs may be cheaper than sampling finite projection DPPs?!

Overview

Introduction

DPP basics

Some insights on finite DPPs

Finite projection DPPs

Continuous projection DPPs

Exact sampling from finite projection DPPs

Approximate sampling from finite projection DPPs

Contributions

Zonotope sampling for finite projection DPPs

Transition

Fast sampling from β -ensembles

Conclusion

Summary of contributions

Open questions and perspectives

Convert a random matrix analysis tool to a computational tool

Finite projection DPP

Approximate sampling
Linear programming



ICML, 2017

β -ensembles

Gibbs sampling
Random matrices

eigvals $\begin{bmatrix} \text{orange diagonal} & 0 \\ \text{blue diagonal} & \text{orange diagonal} \\ 0 & \text{blue diagonal} \end{bmatrix}$

Submitted, 2020

DPP sampling

Monte Carlo integration

Exact sampling
Random linear system



NeurIPS, 2019

Python toolbox

DPPy  
Reproducible research

```
from dppy import *  
# [...]  
dpp.sample()
```

JMLR-MLOSS, 2019

Definition (β -ensemble)

Let (x_1, \dots, x_N) with distribution proportional to

$$\left| \prod_{i < j} (x_j - x_i) \right|^\beta \prod_{n=1}^N e^{-V(x_n)} dx_n,$$

then $\mathcal{X} = \{x_1, \dots, x_N\}$ is called a β -ensemble with potential V .

- ▶ Repulsion characterized by $\prod_{i < j} (x_j - x_i) = \det [x_j^{i-1}]_{i,j=1}^N$.
- ▶ Strength of the repulsion parametrized by $\beta > 0$ (inverse temperature).

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Example ($\beta = 2$, corresponds to a projection DPP)

- ▶ $\mu(dx) = e^{-V(x)} dx$.
- ▶ $K(x, y) = \sum_{k=0}^{N-1} p_k(x) p_k(y)$, $p_k, p_\ell \perp$ polynomials w.r.t. μ .

Classical β -ensembles and random matrix models

Name	Potential $V(x)$	Support
Hermite	$\frac{1}{2\sigma^2}(x - \mu)^2$	\mathbb{R}
Laguerre	$-(k - 1)\log(x) + \frac{1}{\theta}x$	$]0, \infty[$
Jacobi	$-(a - 1)\log(x) - (b - 1)\log(1 - x)$	$]0, 1[$

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$\beta(= 1, 2, 4)$ -ensembles as the eigenvalue distribution of random matrices.

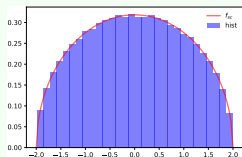
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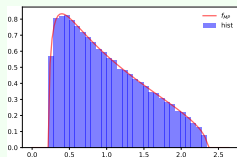
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Example ($\beta = 2$ and $X \sim$ standard complex Gaussian matrix)

- ▶ $X \in \mathbb{C}^{N \times N}$
- ▶ $\text{eigvals}(X + X^H) \sim$ Hermite



- ▶ $X \in \mathbb{C}^{N \times M}$
- ▶ $\text{eigvals}(XX^H) \sim$ Laguerre



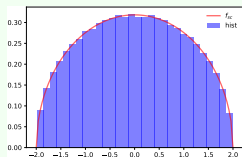
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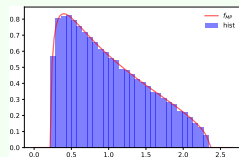
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Random matrix models grant $\mathcal{O}(N^3)$ exact samplers!

Classical β -ensembles and random tridiagonal models

Dumitriu and Edelman (2002) and Killip and Nenciu (2004) derived equivalent random tridiagonal models, valid for $\beta > 0$,

$$\text{eigvals} \left[\begin{array}{cccc} a_1 & \sqrt{b_1} & & (0) \\ \sqrt{b_1} & a_2 & \ddots & \\ & \ddots & \ddots & \sqrt{b_{N-1}} \\ (0) & & \sqrt{b_{N-1}} & a_N \end{array} \right].$$

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Extend tridiagonal models to more general β -ensembles ?

How to randomize the entries of the tridiagonal matrix?

Proposition (Krishnapur et. al, 2016)

Consider the random tridiagonal matrix $J_{\mathbf{a},\mathbf{b}}$ where the entries have joint density

$$\propto e^{-\text{Tr } V(J_{\mathbf{a},\mathbf{b}})} \prod_{n=1}^{N-1} b_n^{\frac{\beta}{2}(N-n)-1}.$$

Then, the eigenvalues of $J_{\mathbf{a},\mathbf{b}}$ have joint density

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- ▶ Perform empirical study of tridiagonal models for polynomial potential V .

Tridiagonal models for polynomial potentials V

When degree $V = 2$,

- ▶ $(a_n), (b_n)$ are independent 😊
- ▶ have easy-to-sample distribution 😊

Example ($V(x) = \frac{1}{2\sigma^2}(x - \mu)^2$)

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$$\text{Example } (V(x) = g_4 x^4 + g_2 x^2)$$

$$a_n \mid \mathbf{a}_{\setminus n}, \mathbf{b}$$

$$\sim \exp\left[-\left(g_4 a_n^4 + a_n^2 [g_2 + 4g_4(b_{n-1} + b_n)] + 4g_4 a_n(a_{n-1}b_{n-1} + a_{n+1}b_n)\right)\right],$$

$$b_n \mid \mathbf{a}, \mathbf{b}_{\setminus n}$$

$$\sim b_n^{\frac{\beta}{2}(N-n)-1} \exp\left[-2\left(g_4 b_n^2 + b_n \left[g_2 + 2g_4(a_n^2 + a_n a_{n+1} + a_{n+1}^2 + b_{n-1} + b_{n+1})\right]\right)\right].$$

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This suggests a Gibbs sampling strategy!

Combining tridiagonal models with Gibbs sampling

Target: β -ensembles with potentials of the form

$$V(x) = g_6x^6 + g_5x^5 + g_4x^4 + g_3x^3 + g_2x^2 + g_1x.$$

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- ▶ Systematic scan Gibbs sampler

for $t = 1$ **to** T

for $n = 1$ **to** N

sample a_n | $\mathbf{a}_{\setminus n}, \mathbf{b}$

sample b_n | $\mathbf{a}, \mathbf{b}_{\setminus n}$ **if** $n < N$

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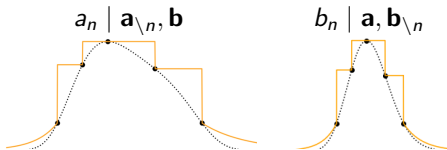
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How does it perform?

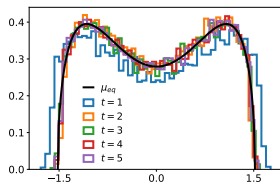
Monitoring of the empirical convergence

Convergence of the empirical marginal distribution to the equilibrium measure.

$$\widehat{\mu}_N^t = \frac{1}{N} \sum_{n=1}^N \delta_{x_n^t} \xrightarrow{N, t \rightarrow \infty} \mu_{\text{eq}}$$

- ▶ $V(x) = \frac{1}{4}x^4$, exact sampling of the conditionals.

$N = 20$

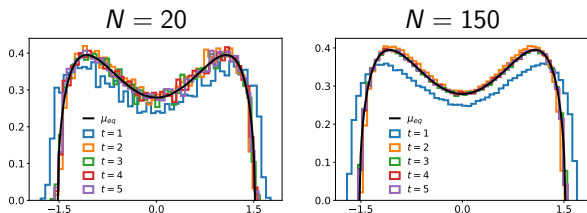


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- ▶ $V(x) = \frac{1}{4}x^4$, exact sampling of the conditionals.

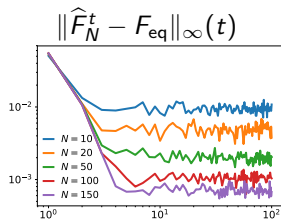
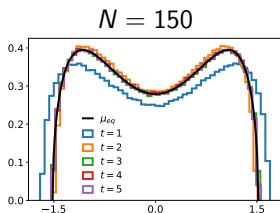
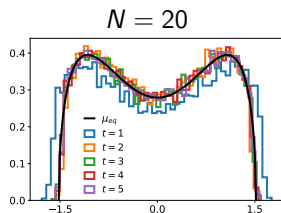


Monitoring of the empirical convergence

Convergence of the empirical marginal distribution to the equilibrium measure.

$$\hat{\mu}_N^t = \frac{1}{N} \sum_{n=1}^N \delta_{x_n^t} \xrightarrow{N, t \rightarrow \infty} \mu_{\text{eq}}.$$

- $V(x) = \frac{1}{4}x^4$, exact sampling of the conditionals.

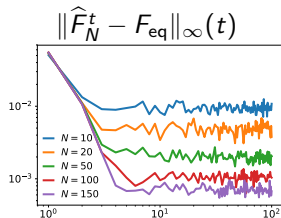
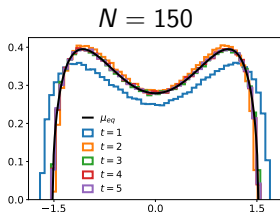
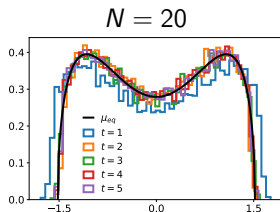


Monitoring of the empirical convergence

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$$\hat{\mu}_N^t = \frac{1}{N} \sum_{n=1}^N \delta_{x_n^t} \xrightarrow{N, t \rightarrow \infty} \mu_{\text{eq}}.$$

- ▶ $V(x) = \frac{1}{4}x^4$, exact sampling of the conditionals.



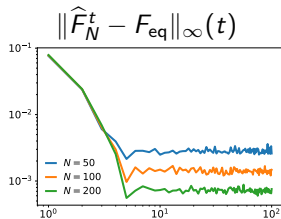
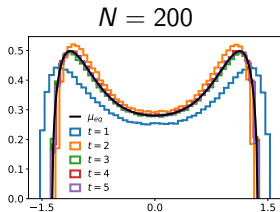
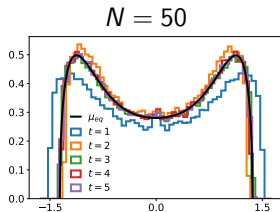
- ▶ Good adequation with the theory.
- ▶ Empirical convergence within $t \leq 10$ Gibbs passes, **only!**

Monitoring of the empirical convergence

Convergence of the empirical marginal distribution to the equilibrium measure.

$$\hat{\mu}_N^t = \frac{1}{N} \sum_{n=1}^N \delta_{x_n^t} \xrightarrow{N, t \rightarrow \infty} \mu_{\text{eq}}.$$

- ▶ $V(x) = \frac{1}{6}x^6$, **approximate** sampling of the conditionals.



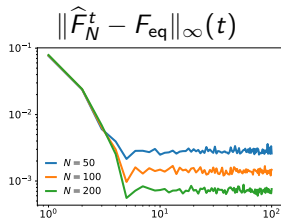
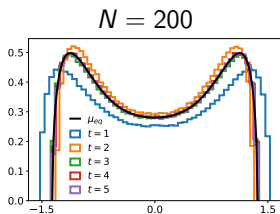
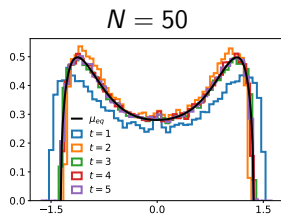
- ▶ Good adequation with the theory.
- ▶ Empirical convergence within $t \leq 10$ Gibbs passes, **only!**

Monitoring of the empirical convergence

Convergence of the empirical marginal distribution to the equilibrium measure.

$$\hat{\mu}_N^t = \frac{1}{N} \sum_{n=1}^N \delta_{x_n^t} \xrightarrow{N, t \rightarrow \infty} \mu_{\text{eq}}.$$

- ▶ $V(x) = \frac{1}{6}x^6$, **approximate** sampling of the conditionals.



- ▶ Good adequation with the theory.
- ▶ Empirical convergence within $t \leq 10$ Gibbs passes, **only!**

Supports the $\mathcal{O}(\log(N))$ mixing time conjecture of Krishnapur et. al (2016).

Overview

Introduction

DPP basics

- Some insights on finite DPPs

- Finite projection DPPs

- Continuous projection DPPs

- Exact sampling from finite projection DPPs

- Approximate sampling from finite projection DPPs

Contributions

- Zonotope sampling for finite projection DPPs

- Transition

- Fast sampling from β -ensembles

Conclusion

- Summary of contributions

- Open questions and perspectives

My Ph.D. in a nutshell

Finite setting

Finite projection DPP
Approximate sampling
Linear programming



ICML, 2017

Continuous setting

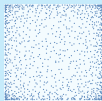
β -ensembles
Gibbs sampling
Random matrices

eigvals $\begin{bmatrix} \text{orange diagonal} & 0 \\ \text{blue diagonal} & \text{orange diagonal} \\ 0 & \text{blue diagonal} \end{bmatrix}$

Submitted, 2020

Continuous setting

Monte Carlo integration
Exact sampling
Random linear system



NeurIPS, 2019

Software engineering

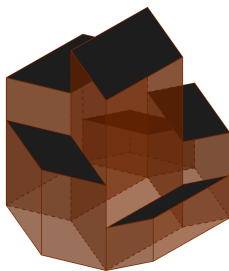
Python toolbox
DPPy 🔄 📄
Reproducible research

```
from dppy import *  
# [...]  
dpp.sample()
```

JMLR-MLOSS, 2019

DPP sampling

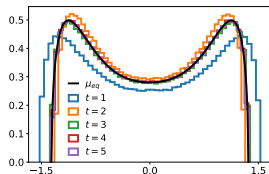
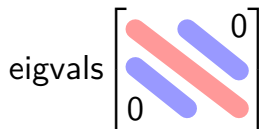
Zonotope sampling for finite projection DPPs



- ▶ New perspective on finite projection DPPs.
- ▶ Combination of geometry, Markov chains and linear programming.
- ▶ Approximate sampler involving randomized linear programs.
- ▶ More efficient exploration of the state space.

ICML, 2017

Tridiagonal models for sampling β -ensembles



- ▶ Unified treatment of tridiagonal models for the classical β -ensembles.
- ▶ Combination of a Gibbs sampler with calculation of eigenvalues.
- ▶ Very fast empirical convergence supporting the $\mathcal{O}(\log(N))$ mixing time conjecture.

Submitted to an international journal, 2020

Monte Carlo integration with DPPs

Let $\{x_1, \dots, x_N\} \sim \text{DPP}(K, \mu)$, where $K(x, y) = \sum_{k=0}^{N-1} \phi_k(x)\phi_k(y)$.

$$\int f(x)\mu(dx) \approx \sum_{n=1}^N \omega_n f(x_n),$$

- ▶ Shed light on the estimator of Ermakov and Zolotukhin (1960)
 - ▶ involving a randomized linear system
 - ▶ provide new simple proofs of its properties

$$\text{Var} = \|f\|^2 - \sum_{k=0}^{N-1} \langle f, \phi_k \rangle^2.$$

- ▶ Numerical comparison with the estimator of Bardenet and Hardy (2020)
- ▶ Tailored implementation of the chain rule.

Adapt the kernel K to the basis where f has a smooth/sparse expansion.

NeurIPS, 2019

DPPy: DPP sampling with Python

📄 [guilgautier / DPPy](#)

📖 Used by 8

👁️ Unwatch 12

★ Unstar 94

🍴 Fork 24

```
from dppy import *  
# [...]  
dpp.sample()
```

- ▶ Open source toolbox 🔄.
- ▶ Implementation of exact and approximate samplers.
- ▶ Extensive documentation 📖.

JMLR-MLOSS, 2019

Overview

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Summary of contributions

Open questions and perspectives

Open questions

- ▶ Zonotope
 - ▶ prove a bound on the mixing time.
 - ▶ extend the LP idea for continuous DPPs.
- ▶ β -ensembles
 - ▶ prove the $\mathcal{O}(\log(N))$ mixing time for the Gibbs sampler.
 - ▶ extend tridiagonal models for multivariate β -ensembles.
- ▶ Efficient sampler for continuous projection DPPs ($d > 1$)?
- ▶ Avoid kernel eigendecomposition for sampling non-projection DPPs?

Perspectives

- ▶ Find a good reparametrization of DPPs where
 - ▶ complex interaction structure vanishes.
 - ▶ efficient sampling can be performed.
- ▶ Continuous extension of sampling by solving linear programs.
- ▶ Sampling by coupling the target DPP with another process.
 - ▶ Decreusefond, Flint, and Low (2013), Launay, Galerne, and Desolneux (2018), and Dereziński, Calandriello, and Valko (2019).
- ▶ Continue developing the DPPy toolbox  .

Thank you!
Ευχαριστώ!
Merci!

My PhD story

COMMENT GÉNÉRER DES RECOMMANDATIONS DIVERSES

From dppy import *

$\det K_S$

NOUVEAUX ALGORITHMES

CONFÉRENCES

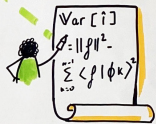
ARTICLES SCIENTIFIQUES

RECHERCHE REPRODUCTIBLE

Guillaume

MISSIONS

Monsieur Jeanne



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L -ensembles and k -DPPs

Definition (L -ensemble)

Let $\mathbf{L} \succeq 0$. The point process defined by

$$\mathbb{P}[\mathcal{X} = S] = \frac{\det \mathbf{L}_S}{\det(I + \mathbf{L})},$$

is called an L -ensemble. It is a DPP with kernel $\mathbf{K} = \mathbf{L}(I + \mathbf{L})^{-1}$.

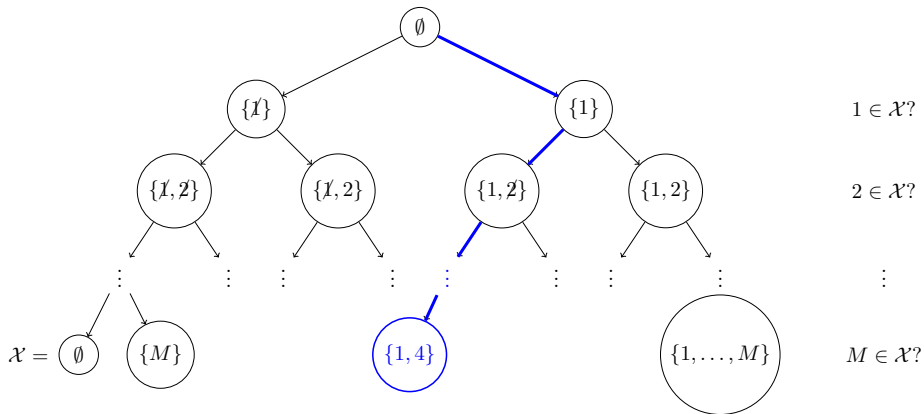
Definition (k -DPP)

Let $\mathbf{L} \succeq 0$ and $k \in \mathbb{N}^*$. The point process defined by

$$\mathbb{P}[\mathcal{X} = S] \propto \det \mathbf{L}_S \mathbb{1}_{|S|=k}.$$

is called a k -DPP.

Chain rule on sets



BH estimator and the multivariate Jacobi ensemble

Natural unbiased estimator of $\int_{\mathbb{X}} f(x)\mu(dx)$

$$\widehat{I}_N^{\text{BH}}(f) = \sum_{n=1}^N \frac{f(x_n)}{K(x_n, x_n)}$$

- ▶ Bardenet and Hardy (2020) show fast CLT, for f essentially \mathcal{C}^1

$$\sqrt{N^{1+1/d}} \left(\widehat{I}_N^{\text{BH}}(f) - \int_{[-1,1]^d} f(x)\omega(x)dx \right) \xrightarrow[N \rightarrow \infty]{\text{law}} \mathcal{N}(0, \mathbf{\Omega}_{f,\omega}^2),$$

with $\mathbf{\Omega}_{f,\omega}^2 \triangleq \frac{1}{2} \sum_{k \in \mathbb{N}^d} (k_1 + \dots + k_d) \mathcal{F} \left[\frac{f\omega}{\omega_{\text{eq}}} \right] (k)^2$

Theorem (Ermakov and Zolotukhin, 1960)

$$f = \sum_{\ell=0}^{M-1} \langle f, \phi_{\ell} \rangle \phi_{\ell}, \quad M \in \mathbb{N} \cup \{\infty\}$$

1. Sample $\{x_1, \dots, x_N\} \sim \text{DPP}(\mu, K)$ with $K(x, y) = \sum_{k=0}^{N-1} \phi_k(x) \phi_k(y)$
2. Random linear system

$$\begin{bmatrix} \phi_0(x_1) & \dots & \phi_{N-1}(x_1) \\ \vdots & & \vdots \\ \phi_0(x_N) & \dots & \phi_{N-1}(x_N) \end{bmatrix} \begin{bmatrix} y_0 \\ \vdots \\ y_{N-1} \end{bmatrix} = \begin{bmatrix} f(x_1) \\ \vdots \\ f(x_N) \end{bmatrix}$$

- ▶ $\mathbb{E}[y_k] = \langle f, \phi_k \rangle = \int f(x) \phi_k(x) \mu(dx)$
- ▶ $\text{Var}[y_k] = \|f\|^2 - \sum_{\ell=0}^{N-1} \langle f, \phi_{\ell} \rangle^2 = \sum_{\ell=N}^{M-1} \langle f, \phi_{\ell} \rangle^2 = 0$ if $M \leq N$
- ▶ $\text{Cov}[y_j, y_k] = 0, j \neq k$

Ermakov and Zolotukhin (1960) estimator

For constant ϕ_0 , e.g., multivariate Jacobi ensemble,

$$\mathbb{E}[y_0] = \phi_0 \int_{\mathbb{X}} f(x) \mu(dx)$$

A direct application of EZ theorem yields

$$\hat{I}_N^{\text{EZ}}(f) \triangleq \frac{y_0}{\phi_0} = \sqrt{\mu([-1, 1]^d)} \frac{\det \Phi_{\phi_0, f}(x_{1:N})}{\det \Phi(x_{1:N})}$$

as an unbiased estimator of $\int f(x) \mu(dx)$

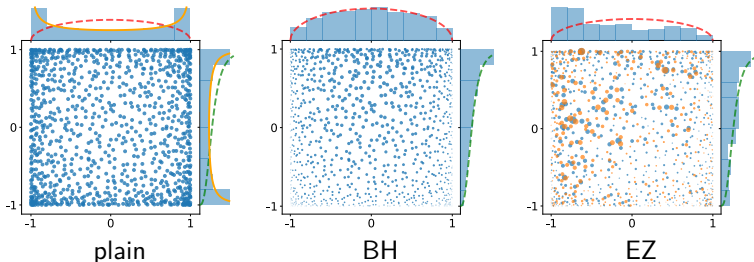
Using $\|\phi_0\| = 1$ and Cramer's rule

$$\Phi_{\phi_0, f} = \begin{bmatrix} f(x_1) & \dots & \psi_{N-1}(x_1) \\ \vdots & & \vdots \\ f(x_N) & \dots & \psi_{N-1}(x_N) \end{bmatrix} \quad \Phi = \begin{bmatrix} \phi_0(x_1) & \dots & \phi_{N-1}(x_1) \\ \vdots & & \vdots \\ \phi_0(x_N) & \dots & \phi_{N-1}(x_N) \end{bmatrix}$$

Comparison weights ω_n BH-EZ

$$\int_{\mathbb{X}} f(x) \mu(dx) \approx \hat{I}_N = \sum_{n=1}^N \omega_n(x_1, \dots, x_N) f(x_n)$$

- weights ω_n



- Non-asymptotic variance

$$\text{Var}[\hat{I}_N^{\text{BH}}] = \frac{1}{2} \int_{\mathbb{X}^2} \left(\frac{f(x)}{K(x,x)} - \frac{f(y)}{K(y,y)} \right)^2 K(x,y)^2 \mu(dx) \mu(dy)$$

$$\text{Var}[\hat{I}_N^{\text{EZ}}] = \|f\|^2 - \sum_{\ell=0}^{N-1} \langle f, \phi_\ell \rangle^2$$

Timings

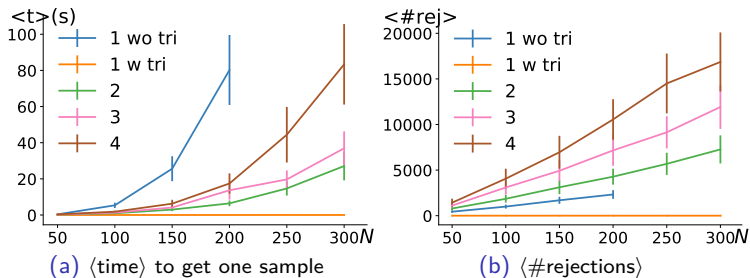


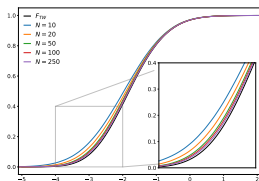
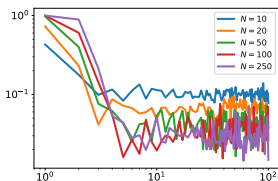
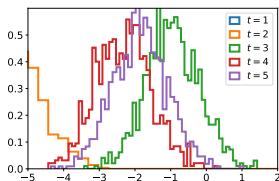
Figure 1: The colors and numbers correspond to the dimension. $a_i, b_i = -1/2$. For $d = 1$, the tridiagonal model (tri) of Killip and Nenciu (BH, 2004) offers tremendous savings, without it is cheaper to get a sample in larger dimension. The number of rejections grows as $N \log(N) 2^d$.

Monitoring of the empirical convergence ($\lambda_{\max}, \beta = 2$)

Convergence of the distribution of the largest eigenvalue to Tracy-Widom.

$$\text{rescaled } \lambda_{\max}^t \xrightarrow[N, t \rightarrow \infty]{\text{law}} \text{TW}_2.$$

- ▶ $V(x) = \frac{1}{4}x^4$, #independent runs = 10^3 .



- ▶ $V(x) = \frac{1}{6}x^6$, #independent runs = 10^3 .

